

# Semi-empirical Likelihood Confidence Intervals for the Differences of Two Populations Based on Fractional Imputation\*

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**Abstract:** Suppose that there are two populations  $x$  and  $y$  with missing data on both of them, where  $x$  has a distribution function  $F(\cdot)$  which is unknown and  $y$  has a distribution function  $G_\theta(\cdot)$  with a probability density function  $g_\theta(\cdot)$  with known form depending on some unknown parameter  $\theta$ . Fractional imputation is used to fill in missing data. The asymptotic distributions of the semi-empirical likelihood ratio statistic are obtained under some mild conditions. Then, empirical likelihood confidence intervals on the differences of  $x$  and  $y$  are constructed.

**Key words:** empirical likelihood, confidence intervals, fractional imputation, missing data

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## 1 Introduction

Missing data are common in opinion polls, market research surveys, medical studies and other scientific experiments. In this situation, the usual inference procedure cannot be applied directly. A common method for handling incomplete data is to impute a value for each missing variable and then apply usual statistical methods to the “complete data” as if they were true observations. Missing data analysis covers a variety of problems that are often seen in practical applications (see [1]). Owen<sup>[2]–[5]</sup> first put forward the technique of empirical likelihood in nonparametric statistics. Recently, Wang and Rao<sup>[6]–[7]</sup> use an empirical likelihood method to construct confidence intervals for the response means in linear

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and nonparametric models. In this paper, we focus on constructing confidence intervals for various differences of two populations  $x$  and  $y$  which have distribution functions  $F(\cdot)$  and  $G_\theta(\cdot)$ , where  $F(\cdot)$  is unknown and  $G_\theta(\cdot)$  with probability density function  $g_\theta(\cdot)$  is of known form depending on some unknown parameter  $\theta$ . Let  $\Delta$  denote the differences of  $x$  and  $y$  such as the differences of the means and the distribution functions of two populations. This model is often seen in practical applications. For example, doctors intend to use a new medicine  $A$  to cure a specified illness.  $B$  is also used to treat the disease.  $B$  is known very well and  $A$  is less known by doctors. Many clinic experiments should be done to obtain the data of curative effects between  $A$  and  $B$ . We are interested in comparing noticeable differences between  $A$  and  $B$ . We take  $B$  and  $A$  as  $x$  and  $y$ , respectively. Thus, we want to know the differences  $\Delta$  of  $x$  and  $y$ , such as

$$\Delta = Ex - Ey, \quad \Delta = P(x \leq y), \quad \Delta = F(x_0) - G_{\theta_0}(x_0) \quad (x_0 \text{ is fixed})$$

and so on. In this paper, we suppose the following information is available

$$E_F \omega(x, \theta, \Delta) = 0, \tag{1.1}$$

where  $\omega$  is a function of known form.

In this paper, we use the information (1.1) to construct empirical likelihood confidence intervals on  $\Delta$ . Thus, we can test the hypotheses on the differences  $\Delta$  of  $x$  and  $y$ . We suppose that the hypothesis is

$$H_0 : \Delta = \Delta_0 \text{ for some known } \Delta_0.$$

If we wish to test the hypothesis that there is no noticeable differences between  $x$  and  $y$ , we let  $\Delta_0 = 0$ . If  $\Delta_0$  is in the above interval, we accept the hypothesis; otherwise, the hypothesis should be rejected.

Consider the following simple random samples of incomplete data associated with populations  $(x, \delta_x)$  and  $(y, \delta_y)$ ,

$$(x_i, \delta_{x_i}), \quad i = 1, \dots, m, \quad (y_j, \delta_{y_j}), \quad j = 1, \dots, n,$$

where

$$\delta_{x_i} = \begin{cases} 0, & \text{if } x_i \text{ is missing;} \\ 1, & \text{else,} \end{cases}$$

$$\delta_{y_j} = \begin{cases} 0, & \text{if } y_j \text{ is missing;} \\ 1, & \text{else.} \end{cases}$$

Suppose that  $x$  and  $y$  are missing completely at random (MCAR)

$$P(\delta_{x_i} = 1|x) = P(\delta_{x_i} = 1) = P_1(\text{constant}),$$

$$P(\delta_{y_j} = 1|y) = P(\delta_{y_j} = 1) = P_2(\text{constant}).$$

We also assume that  $(x, \delta_{x_i})$  and  $(y, \delta_{y_j})$  are independent. Let

$$r_x = \sum_{i=1}^m \delta_{x_i}, \quad m_x = m - r_x, \quad r_y = \sum_{j=1}^n \delta_{y_j}, \quad m_y = n - r_y,$$

$$s_{r_x} = \{i : \delta_{x_i} = 1, i = 1, \dots, m\}, \quad s_{m_x} = \{i : \delta_{x_i} = 0, i = 1, \dots, m\},$$

$$s_{r_y} = \{j : \delta_{y_j} = 1, j = 1, \dots, n\}, \quad s_{m_y} = \{j : \delta_{y_j} = 0, j = 1, \dots, n\}.$$